

Package: extraDistr (via r-universe)

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Type Package

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Description Density, distribution function, quantile function and random generation for a number of univariate and multivariate distributions. This package implements the following distributions: Bernoulli, beta-binomial, beta-negative binomial, beta prime, Bhattacharjee, Birnbaum-Saunders, bivariate normal, bivariate Poisson, categorical, Dirichlet, Dirichlet-multinomial, discrete gamma, discrete Laplace, discrete normal, discrete uniform, discrete Weibull, Frechet, gamma-Poisson, generalized extreme value, Gompertz, generalized Pareto, Gumbel, half-Cauchy, half-normal, half-t, Huber density, inverse chi-squared, inverse-gamma, Kumaraswamy, Laplace, location-scale t, logarithmic, Lomax, multivariate hypergeometric, multinomial, negative hypergeometric, non-standard beta, normal mixture, Poisson mixture, Pareto, power, reparametrized beta, Rayleigh, shifted Gompertz, Skellam, slash, triangular, truncated binomial, truncated normal, truncated Poisson, Tukey lambda, Wald, zero-inflated binomial, zero-inflated negative binomial, zero-inflated Poisson.

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Bernoulli	<i>Bernoulli distribution</i>
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Description

Probability mass function, distribution function, quantile function and random generation for the Bernoulli distribution.

Usage

```
dbern(x, prob = 0.5, log = FALSE)
```

```
pbern(q, prob = 0.5, lower.tail = TRUE, log.p = FALSE)
```

```
qbern(p, prob = 0.5, lower.tail = TRUE, log.p = FALSE)
```

```
rbern(n, prob = 0.5)
```

Arguments

x, q	vector of quantiles.
prob	probability of success; ($0 < \text{prob} < 1$).
log, log.p	logical; if TRUE, probabilities p are given as $\log(p)$.
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.

n number of observations. If `length(n) > 1`, the length is taken to be the number required.

See Also

[Binomial](#)

Examples

```
prop.table(table(rbern(1e5, 0.5)))
```

BetaBinom

Beta-binomial distribution

Description

Probability mass function and random generation for the beta-binomial distribution.

Usage

```
dbbinom(x, size, alpha = 1, beta = 1, log = FALSE)
```

```
pbbinom(q, size, alpha = 1, beta = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rbbinom(n, size, alpha = 1, beta = 1)
```

Arguments

x, q vector of quantiles.
 size number of trials (zero or more).
 alpha, beta non-negative parameters of the beta distribution.
 log, log.p logical; if TRUE, probabilities p are given as log(p).
 lower.tail logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
 n number of observations. If `length(n) > 1`, the length is taken to be the number required.

Details

If $p \sim \text{Beta}(\alpha, \beta)$ and $X \sim \text{Binomial}(n, p)$, then $X \sim \text{BetaBinomial}(n, \alpha, \beta)$.

Probability mass function

$$f(x) = \binom{n}{x} \frac{B(x + \alpha, n - x + \beta)}{B(\alpha, \beta)}$$

Cumulative distribution function is calculated using recursive algorithm that employs the fact that $\Gamma(x) = (x - 1)!$, and $B(x, y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}$, and that $\binom{n}{k} = \prod_{i=1}^k \frac{n+1-i}{i}$. This enables re-writing probability mass function as

$$f(x) = \left(\prod_{i=1}^x \frac{n+1-i}{i} \right) \frac{(\alpha+x-1)! (\beta+n-x-1)!}{(\alpha+\beta+n-1)! B(\alpha, \beta)}$$

what makes recursive updating from x to $x+1$ easy using the properties of factorials

$$f(x+1) = \left(\prod_{i=1}^x \frac{n+1-i}{i} \right) \frac{n+1-x+1}{x+1} \frac{(\alpha+x-1)! (\alpha+x) (\beta+n-x-1)! (\beta+n-x)^{-1}}{(\alpha+\beta+n-1)! (\alpha+\beta+n) B(\alpha, \beta)}$$

and let's us efficiently calculate cumulative distribution function as a sum of probability mass functions

$$F(x) = \sum_{k=0}^x f(k)$$

See Also

[Beta, Binomial](#)

Examples

```
x <- rbbinom(1e5, 1000, 5, 13)
xx <- 0:1000
hist(x, 100, freq = FALSE)
lines(xx-0.5, dbbinom(xx, 1000, 5, 13), col = "red")
hist(pbbinom(x, 1000, 5, 13))
xx <- seq(0, 1000, by = 0.1)
plot(ecdf(x))
lines(xx, pbbinom(xx, 1000, 5, 13), col = "red", lwd = 2)
```

BetaNegBinom

Beta-negative binomial distribution

Description

Probability mass function and random generation for the beta-negative binomial distribution.

Usage

```
dbnbinom(x, size, alpha = 1, beta = 1, log = FALSE)
```

```
pbnbinom(q, size, alpha = 1, beta = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rbbinom(n, size, alpha = 1, beta = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>size</code>	number of trials (zero or more). Must be strictly positive, need not be integer.
<code>alpha, beta</code>	non-negative parameters of the beta distribution.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If $p \sim \text{Beta}(\alpha, \beta)$ and $X \sim \text{NegBinomial}(r, p)$, then $X \sim \text{BetaNegBinomial}(r, \alpha, \beta)$.

Probability mass function

$$f(x) = \frac{\Gamma(r+x) B(\alpha+r, \beta+x)}{x! \Gamma(r) B(\alpha, \beta)}$$

Cumulative distribution function is calculated using recursive algorithm that employs the fact that $\Gamma(x) = (x-1)!$ and $B(x, y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}$. This enables re-writing probability mass function as

$$f(x) = \frac{(r+x-1)!}{x! \Gamma(r)} \frac{(\alpha+r-1)! (\beta+x-1)!}{(\alpha+\beta+r+x-1)! B(\alpha, \beta)}$$

what makes recursive updating from x to $x+1$ easy using the properties of factorials

$$f(x+1) = \frac{(r+x-1)! (r+x)}{x! (x+1) \Gamma(r)} \frac{(\alpha+r-1)! (\beta+x-1)! (\beta+x)}{(\alpha+\beta+r+x-1)! (\alpha+\beta+r+x) B(\alpha, \beta)}$$

and let's us efficiently calculate cumulative distribution function as a sum of probability mass functions

$$F(x) = \sum_{k=0}^x f(k)$$

See Also

[Beta](#), [NegBinomial](#)

Examples

```
x <- rbnbinom(1e5, 1000, 5, 13)
xx <- 0:1e5
hist(x, 100, freq = FALSE)
lines(xx-0.5, dbnbinom(xx, 1000, 5, 13), col = "red")
hist(pbnbinom(x, 1000, 5, 13))
xx <- seq(0, 1e5, by = 0.1)
plot(ecdf(x))
lines(xx, pbnbinom(xx, 1000, 5, 13), col = "red", lwd = 2)
```

BetaPrime	<i>Beta prime distribution</i>
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Description

Density, distribution function, quantile function and random generation for the beta prime distribution.

Usage

```
dbetapr(x, shape1, shape2, scale = 1, log = FALSE)
pbetapr(q, shape1, shape2, scale = 1, lower.tail = TRUE, log.p = FALSE)
qbetapr(p, shape1, shape2, scale = 1, lower.tail = TRUE, log.p = FALSE)
rbetapr(n, shape1, shape2, scale = 1)
```

Arguments

x, q	vector of quantiles.
shape1, shape2	non-negative parameters.
scale	positive valued scale parameter.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

If $X \sim \text{Beta}(\alpha, \beta)$, then $\frac{X}{1-X} \sim \text{BetaPrime}(\alpha, \beta)$.

Probability density function

$$f(x) = \frac{(x/\sigma)^{\alpha-1} (1+x/\sigma)^{-\alpha-\beta}}{B(\alpha, \beta)\sigma}$$

Cumulative distribution function

$$F(x) = I_{\frac{x/\sigma}{1+x/\sigma}}(\alpha, \beta)$$

See Also

[Beta](#)

Examples

```
x <- rbetapr(1e5, 5, 3, 2)
hist(x, 350, freq = FALSE, xlim = c(0, 100))
curve/dbetapr(x, 5, 3, 2), 0, 100, col = "red", add = TRUE, n = 500)
hist(pbetapr(x, 5, 3, 2))
plot(ecdf(x), xlim = c(0, 100))
curve(pbetapr(x, 5, 3, 2), 0, 100, col = "red", add = TRUE, n = 500)
```

Bhattacharjee

*Bhattacharjee distribution***Description**

Density, distribution function, and random generation for the Bhattacharjee distribution.

Usage

```
dbhatt(x, mu = 0, sigma = 1, a = sigma, log = FALSE)
pbhatt(q, mu = 0, sigma = 1, a = sigma, lower.tail = TRUE, log.p = FALSE)
rbhatt(n, mu = 0, sigma = 1, a = sigma)
```

Arguments

x, q	vector of quantiles.
mu, sigma, a	location, scale and shape parameters. Scale and shape must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If $Z \sim \text{Normal}(0, 1)$ and $U \sim \text{Uniform}(0, 1)$, then $Z + U$ follows Bhattacharjee distribution.

Probability density function

$$f(z) = \frac{1}{2a} \left[\Phi \left(\frac{x - \mu + a}{\sigma} \right) - \Phi \left(\frac{x - \mu - a}{\sigma} \right) \right]$$

Cumulative distribution function

$$F(z) = \frac{\sigma}{2a} \left[(x - \mu) \Phi \left(\frac{x - \mu + a}{\sigma} \right) - (x - \mu) \Phi \left(\frac{x - \mu - a}{\sigma} \right) + \phi \left(\frac{x - \mu + a}{\sigma} \right) - \phi \left(\frac{x - \mu - a}{\sigma} \right) \right]$$

References

Bhattacharjee, G.P., Pandit, S.N.N., and Mohan, R. (1963). Dimensional chains involving rectangular and normal error-distributions. *Technometrics*, 5, 404-406.

Examples

```
x <- rbhatt(1e5, 5, 3, 5)
hist(x, 100, freq = FALSE)
curve(dbhatt(x, 5, 3, 5), -20, 20, col = "red", add = TRUE)
hist(pbhatt(x, 5, 3, 5))
plot(ecdf(x))
curve(pbhatt(x, 5, 3, 5), -20, 20, col = "red", lwd = 2, add = TRUE)
```

BirbaumSaunders	<i>Birbaum-Saunders (fatigue life) distribution</i>
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Description

Density, distribution function, quantile function and random generation for the Birbaum-Saunders (fatigue life) distribution.

Usage

```
dfatigue(x, alpha, beta = 1, mu = 0, log = FALSE)
pfatigue(q, alpha, beta = 1, mu = 0, lower.tail = TRUE, log.p = FALSE)
qfatigue(p, alpha, beta = 1, mu = 0, lower.tail = TRUE, log.p = FALSE)
rfatigue(n, alpha, beta = 1, mu = 0)
```

Arguments

x, q	vector of quantiles.
alpha, beta, mu	shape, scale and location parameters. Scale and shape must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \left(\frac{\sqrt{\frac{x-\mu}{\beta}} + \sqrt{\frac{\beta}{x-\mu}}}{2\alpha(x-\mu)} \right) \phi \left(\frac{1}{\alpha} \left(\sqrt{\frac{x-\mu}{\beta}} - \sqrt{\frac{\beta}{x-\mu}} \right) \right)$$

Cumulative distribution function

$$F(x) = \Phi \left(\frac{1}{\alpha} \left(\sqrt{\frac{x-\mu}{\beta}} - \sqrt{\frac{\beta}{x-\mu}} \right) \right)$$

Quantile function

$$F^{-1}(p) = \left[\frac{\alpha}{2} \Phi^{-1}(p) + \sqrt{\left(\frac{\alpha}{2} \Phi^{-1}(p) \right)^2 + 1} \right]^2 \beta + \mu$$

References

Birnbaum, Z. W. and Saunders, S. C. (1969). A new family of life distributions. *Journal of Applied Probability*, 6(2), 637-652.

Desmond, A. (1985) Stochastic models of failure in random environments. *Canadian Journal of Statistics*, 13, 171-183.

Vilca-Labra, F., and Leiva-Sanchez, V. (2006). A new fatigue life model based on the family of skew-elliptical distributions. *Communications in Statistics-Theory and Methods*, 35(2), 229-244.

Leiva, V., Sanhueza, A., Sen, P. K., and Paula, G. A. (2008). Random number generators for the generalized Birnbaum-Saunders distribution. *Journal of Statistical Computation and Simulation*, 78(11), 1105-1118.

Examples

```
x <- rfatigue(1e5, .5, 2, 5)
hist(x, 100, freq = FALSE)
curve(dfatigue(x, .5, 2, 5), 2, 20, col = "red", add = TRUE)
hist(pfatigue(x, .5, 2, 5))
plot(ecdf(x))
curve(pfatigue(x, .5, 2, 5), 2, 20, col = "red", lwd = 2, add = TRUE)
```

BivNormal

Bivariate normal distribution

Description

Density, distribution function and random generation for the bivariate normal distribution.

Usage

```
dbvnorm(
  x,
  y = NULL,
  mean1 = 0,
  mean2 = mean1,
  sd1 = 1,
  sd2 = sd1,
  cor = 0,
  log = FALSE
)
```

```
rbvnorm(n, mean1 = 0, mean2 = mean1, sd1 = 1, sd2 = sd1, cor = 0)
```

Arguments

`x, y` vectors of quantiles; alternatively `x` may be a two-column matrix (or `data.frame`) and `y` may be omitted.

`mean1, mean2` vectors of means.

`sd1, sd2` vectors of standard deviations.

`cor` vector of correlations ($-1 < \text{cor} < 1$).

`log` logical; if TRUE, probabilities `p` are given as `log(p)`.

`n` number of observations. If `length(n) > 1`, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{1}{2\pi\sqrt{1-\rho^2}\sigma_1\sigma_2} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left[\left(\frac{x_1 - \mu_1}{\sigma_1} \right)^2 - 2\rho \left(\frac{x_1 - \mu_1}{\sigma_1} \right) \left(\frac{x_2 - \mu_2}{\sigma_2} \right) + \left(\frac{x_2 - \mu_2}{\sigma_2} \right)^2 \right] \right\}$$

References

Krishnamoorthy, K. (2006). Handbook of Statistical Distributions with Applications. Chapman & Hall/CRC

Mukhopadhyay, N. (2000). Probability and statistical inference. Chapman & Hall/CRC

See Also

[Normal](#)

Examples

```
y <- x <- seq(-4, 4, by = 0.25)
z <- outer(x, y, function(x, y) dbvnorm(x, y, cor = -0.75))
persp(x, y, z)
```

```

y <- x <- seq(-4, 4, by = 0.25)
z <- outer(x, y, function(x, y) dbvnorm(x, y, cor = -0.25))
persp(x, y, z)

```

BivPoiss

Bivariate Poisson distribution

Description

Probability mass function and random generation for the bivariate Poisson distribution.

Usage

```
dbvpois(x, y = NULL, a, b, c, log = FALSE)
```

```
rbvpois(n, a, b, c)
```

Arguments

x, y	vectors of quantiles; alternatively x may be a two-column matrix (or data.frame) and y may be omitted.
a, b, c	positive valued parameters.
log	logical; if TRUE, probabilities p are given as log(p).
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

Probability mass function

$$f(x, y) = \exp\{-(a + b + c)\} \frac{a^x b^y}{x! y!} \sum_{k=0}^{\min(x, y)} \binom{x}{k} \binom{y}{k} k! \left(\frac{c}{ab}\right)^k$$

References

- Karlis, D. and Ntzoufras, I. (2003). Analysis of sports data by using bivariate Poisson models. *Journal of the Royal Statistical Society: Series D (The Statistician)*, 52(3), 381-393.
- Kocherlakota, S. and Kocherlakota, K. (1992) *Bivariate Discrete Distributions*. New York: Dekker.
- Johnson, N., Kotz, S. and Balakrishnan, N. (1997). *Discrete Multivariate Distributions*. New York: Wiley.
- Holgate, P. (1964). Estimation for the bivariate Poisson distribution. *Biometrika*, 51(1-2), 241-287.
- Kawamura, K. (1984). Direct calculation of maximum likelihood estimator for the bivariate Poisson distribution. *Kodai mathematical journal*, 7(2), 211-221.

See Also[Poisson](#)**Examples**

```
x <- rbvpois(5000, 7, 8, 5)
image(prop.table(table(x[,1], x[,2])))
colMeans(x)
```

Categorical*Categorical distribution*

Description

Probability mass function, distribution function, quantile function and random generation for the categorical distribution.

Usage

```
dcat(x, prob, log = FALSE)

pcat(q, prob, lower.tail = TRUE, log.p = FALSE)

qcat(p, prob, lower.tail = TRUE, log.p = FALSE, labels)

rcat(n, prob, labels)

rcatlp(n, log_prob, labels)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>prob, log_prob</code>	vector of length m , or m -column matrix of non-negative weights (or their logarithms in <code>log_prob</code>).
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>labels</code>	if provided, labeled factor vector is returned. Number of labels needs to be the same as number of categories (number of columns in <code>prob</code>).
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability mass function

$$\Pr(X = k) = \frac{w_k}{\sum_{j=1}^m w_j}$$

Cumulative distribution function

$$\Pr(X \leq k) = \frac{\sum_{i=1}^k w_i}{\sum_{j=1}^m w_j}$$

It is possible to sample from categorical distribution parametrized by vector of unnormalized log-probabilities $\alpha_1, \dots, \alpha_m$ without leaving the log space by employing the Gumbel-max trick (Maddison, Tarlow and Minka, 2014). If g_1, \dots, g_m are samples from Gumbel distribution with cumulative distribution function $F(g) = \exp(-\exp(-g))$, then $k = \arg \max_i \{g_i + \alpha_i\}$ is a draw from categorical distribution parametrized by vector of probabilities p_1, \dots, p_m , such that $p_i = \exp(\alpha_i) / [\sum_{j=1}^m \exp(\alpha_j)]$. This is implemented in `rcat1p` function parametrized by vector of log-probabilities `log_prob`.

References

Maddison, C. J., Tarlow, D., & Minka, T. (2014). A* sampling. [In:] Advances in Neural Information Processing Systems (pp. 3086-3094). <https://arxiv.org/abs/1411.0030>

Examples

```
# Generating 10 random draws from categorical distribution
# with k=3 categories occurring with equal probabilities
# parametrized using a vector

rcat(10, c(1/3, 1/3, 1/3))

# or with k=5 categories parametrized using a matrix of probabilities
# (generated from Dirichlet distribution)

p <- rdirichlet(10, c(1, 1, 1, 1, 1))
rcat(10, p)

x <- rcat(1e5, c(0.2, 0.4, 0.3, 0.1))
plot(prop.table(table(x)), type = "h")
lines(0:5, dcat(0:5, c(0.2, 0.4, 0.3, 0.1)), col = "red")

p <- rdirichlet(1, rep(1, 20))
x <- rcat(1e5, matrix(rep(p, 2), nrow = 2, byrow = TRUE))
xx <- 0:21
plot(prop.table(table(x)))
lines(xx, dcat(xx, p), col = "red")

xx <- seq(0, 21, by = 0.01)
plot(ecdf(x))
lines(xx, pcat(xx, p), col = "red", lwd = 2)
```

```
pp <- seq(0, 1, by = 0.001)
plot(ecdf(x))
lines(qcat(pp, p), pp, col = "red", lwd = 2)
```

Dirichlet

*Dirichlet distribution***Description**

Density function and random generation for the Dirichlet distribution.

Usage

```
ddirichlet(x, alpha, log = FALSE)

rdirichlet(n, alpha)
```

Arguments

x	<i>k</i> -column matrix of quantiles.
alpha	<i>k</i> -values vector or <i>k</i> -column matrix; concentration parameter. Must be positive.
log	logical; if TRUE, probabilities p are given as log(p).
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{\Gamma(\sum_k \alpha_k)}{\prod_k \Gamma(\alpha_k)} \prod_k x_k^{\alpha_k - 1}$$

References

Devroye, L. (1986). Non-Uniform Random Variate Generation. Springer-Verlag.

Examples

```
# Generating 10 random draws from Dirichlet distribution
# parametrized using a vector

rdirichlet(10, c(1, 1, 1, 1))

# or parametrized using a matrix where each row
# is a vector of parameters

alpha <- matrix(c(1, 1, 1, 1:3, 7:9), ncol = 3, byrow = TRUE)
rdirichlet(10, alpha)
```

DirMnom

*Dirichlet-multinomial (multivariate Polya) distribution***Description**

Density function, cumulative distribution function and random generation for the Dirichlet-multinomial (multivariate Polya) distribution.

Usage

```
ddirmnom(x, size, alpha, log = FALSE)
```

```
rdirnmom(n, size, alpha)
```

Arguments

x	<i>k</i> -column matrix of quantiles.
size	numeric vector; number of trials (zero or more).
alpha	<i>k</i> -values vector or <i>k</i> -column matrix; concentration parameter. Must be positive.
log	logical; if TRUE, probabilities p are given as log(p).
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

If $(p_1, \dots, p_k) \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_k)$ and $(x_1, \dots, x_k) \sim \text{Multinomial}(n, p_1, \dots, p_k)$, then $(x_1, \dots, x_k) \sim \text{DirichletMultinomial}(n, \alpha_1, \dots, \alpha_k)$.

Probability density function

$$f(x) = \frac{(n!) \Gamma(\sum \alpha_k)}{\Gamma(n + \sum \alpha_k)} \prod_{k=1}^K \frac{\Gamma(x_k + \alpha_k)}{(x_k!) \Gamma(\alpha_k)}$$

References

Gentle, J.E. (2006). Random number generation and Monte Carlo methods. Springer.

Kvam, P. and Day, D. (2001) The multivariate Polya distribution in combat modeling. Naval Research Logistics, 48, 1-17.

See Also

[Dirichlet, Multinomial](#)

`DiscreteGamma`*Discrete gamma distribution*

Description

Probability mass function, distribution function and random generation for discrete gamma distribution.

Usage

```
ddgamma(x, shape, rate = 1, scale = 1/rate, log = FALSE)
```

```
pdgamma(q, shape, rate = 1, scale = 1/rate, lower.tail = TRUE, log.p = FALSE)
```

```
rdgamma(n, shape, rate = 1, scale = 1/rate)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>shape, scale</code>	shape and scale parameters. Must be positive, scale strictly.
<code>rate</code>	an alternative way to specify the scale.
<code>log, log.p</code>	logical; if TRUE, probabilities p are given as log(p).
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Probability mass function of discrete gamma distribution $f_Y(y)$ is defined by discretization of continuous gamma distribution $f_Y(y) = S_X(y) - S_X(y + 1)$ where S_X is a survival function of continuous gamma distribution.

References

Chakraborty, S. and Chakravarty, D. (2012). Discrete Gamma distributions: Properties and parameter estimations. *Communications in Statistics-Theory and Methods*, 41(18), 3301-3324.

See Also

[GammaDist](#), [DiscreteNormal](#)

Examples

```
x <- rdgamma(1e5, 9, 1)
xx <- 0:50
plot(prop.table(table(x)))
lines(xx, dgamma(xx, 9, 1), col = "red")
hist(pdgamma(x, 9, 1))
plot(ecdf(x))
xx <- seq(0, 50, 0.1)
lines(xx, pdgamma(xx, 9, 1), col = "red", lwd = 2, type = "s")
```

DiscreteLaplace

Discrete Laplace distribution

Description

Probability mass, distribution function and random generation for the discrete Laplace distribution parametrized by location and scale.

Usage

```
ddlplace(x, location, scale, log = FALSE)

pdlplace(q, location, scale, lower.tail = TRUE, log.p = FALSE)

rdlplace(n, location, scale)
```

Arguments

x, q	vector of quantiles.
location	location parameter.
scale	scale parameter; $0 < \text{scale} < 1$.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If $U \sim \text{Geometric}(1 - p)$ and $V \sim \text{Geometric}(1 - p)$, then $U - V \sim \text{DiscreteLaplace}(p)$, where geometric distribution is related to discrete Laplace distribution in similar way as exponential distribution is related to Laplace distribution.

Probability mass function

$$f(x) = \frac{1 - p}{1 + p} p^{|x - \mu|}$$

Cumulative distribution function

$$F(x) = \begin{cases} \frac{p^{-|x-\mu|}}{1+p} & x < 0 \\ 1 - \frac{p^{|x-\mu|+1}}{1+p} & x \geq 0 \end{cases}$$

References

Inusah, S., & Kozubowski, T.J. (2006). A discrete analogue of the Laplace distribution. *Journal of statistical planning and inference*, 136(3), 1090-1102.

Kotz, S., Kozubowski, T., & Podgorski, K. (2012). *The Laplace distribution and generalizations: a revisit with applications to communications, economics, engineering, and finance*. Springer Science & Business Media.

Examples

```
p <- 0.45
x <- rdlaplace(1e5, 0, p)
xx <- seq(-200, 200, by = 1)
plot(prop.table(table(x)))
lines(xx, ddlaplace(xx, 0, p), col = "red")
hist(pdlaplace(x, 0, p))
plot(ecdf(x))
lines(xx, pdlaplace(xx, 0, p), col = "red", type = "s")
```

DiscreteNormal

Discrete normal distribution

Description

Probability mass function, distribution function and random generation for discrete normal distribution.

Usage

```
ddnorm(x, mean = 0, sd = 1, log = FALSE)

pdnorm(q, mean = 0, sd = 1, lower.tail = TRUE, log.p = FALSE)

rdnorm(n, mean = 0, sd = 1)
```

Arguments

x, q	vector of quantiles.
mean	vector of means.
sd	vector of standard deviations.

<code>log, log.p</code>	logical; if TRUE, probabilities p are given as log(p).
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Probability mass function

$$f(x) = \Phi\left(\frac{x - \mu + 1}{\sigma}\right) - \Phi\left(\frac{x - \mu}{\sigma}\right)$$

Cumulative distribution function

$$F(x) = \Phi\left(\frac{\lfloor x \rfloor + 1 - \mu}{\sigma}\right)$$

References

Roy, D. (2003). The discrete normal distribution. *Communications in Statistics-Theory and Methods*, 32, 1871-1883.

See Also

[Normal](#)

Examples

```
x <- rdnorm(1e5, 0, 3)
xx <- -15:15
plot(prop.table(table(x)))
lines(xx, ddnorm(xx, 0, 3), col = "red")
hist(pdnorm(x, 0, 3))
plot(ecdf(x))
xx <- seq(-15, 15, 0.1)
lines(xx, pdnorm(xx, 0, 3), col = "red", lwd = 2, type = "s")
```

DiscreteUniform

Discrete uniform distribution

Description

Probability mass function, distribution function, quantile function and random generation for the discrete uniform distribution.

Usage

```

ddunif(x, min, max, log = FALSE)

pdunif(q, min, max, lower.tail = TRUE, log.p = FALSE)

qdunif(p, min, max, lower.tail = TRUE, log.p = FALSE)

rdunif(n, min, max)

```

Arguments

<code>x, q</code>	vector of quantiles.
<code>min, max</code>	lower and upper limits of the distribution. Must be finite.
<code>log, log.p</code>	logical; if TRUE, probabilities p are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If $\text{min} == \text{max}$, then discrete uniform distribution is a degenerate distribution.

Examples

```

x <- rdunif(1e5, 1, 10)
xx <- -1:11
plot(prop.table(table(x)), type = "h")
lines(xx, ddunif(xx, 1, 10), col = "red")
hist(pdunif(x, 1, 10))
xx <- seq(-1, 11, by = 0.01)
plot(ecdf(x))
lines(xx, pdunif(xx, 1, 10), col = "red")

```

DiscreteWeibull

Discrete Weibull distribution (type I)

Description

Density, distribution function, quantile function and random generation for the discrete Weibull (type I) distribution.

Usage

```

ddweibull(x, shape1, shape2, log = FALSE)

pdweibull(q, shape1, shape2, lower.tail = TRUE, log.p = FALSE)

qdweibull(p, shape1, shape2, lower.tail = TRUE, log.p = FALSE)

rdweibull(n, shape1, shape2)

```

Arguments

x, q	vector of quantiles.
shape1, shape2	parameters (named q, β). Values of shape2 need to be positive and $0 < \text{shape1} < 1$.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability mass function

$$f(x) = q^{x^\beta} - q^{(x+1)^\beta}$$

Cumulative distribution function

$$F(x) = 1 - q^{(x+1)^\beta}$$

Quantile function

$$F^{-1}(p) = \left\lceil \left(\frac{\log(1-p)}{\log(q)} \right)^{1/\beta} - 1 \right\rceil$$

References

Nakagawa, T. and Osaki, S. (1975). The Discrete Weibull Distribution. IEEE Transactions on Reliability, R-24, 300-301.

Kulasekera, K.B. (1994). Approximate MLE's of the parameters of a discrete Weibull distribution with type I censored data. Microelectronics Reliability, 34(7), 1185-1188.

Khan, M.A., Khalique, A. and Abouammoh, A.M. (1989). On estimating parameters in a discrete Weibull distribution. IEEE Transactions on Reliability, 38(3), 348-350.

See Also

[Weibull](#)

Examples

```
x <- rdweibull(1e5, 0.32, 1)
xx <- seq(-2, 100, by = 1)
plot(prop.table(table(x)), type = "h")
lines(xx, ddweibull(xx, .32, 1), col = "red")

# Notice: distribution of F(X) is far from uniform:
hist(pdweibull(x, .32, 1), 50)

plot(ecdf(x))
lines(xx, pdweibull(xx, .32, 1), col = "red", lwd = 2, type = "s")
```

 Frechet

Frechet distribution

Description

Density, distribution function, quantile function and random generation for the Frechet distribution.

Usage

```
dfrechet(x, lambda = 1, mu = 0, sigma = 1, log = FALSE)
pfrechet(q, lambda = 1, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)
qfrechet(p, lambda = 1, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)
rfrechet(n, lambda = 1, mu = 0, sigma = 1)
```

Arguments

x, q	vector of quantiles.
lambda, sigma, mu	shape, scale, and location parameters. Scale and shape must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{\lambda}{\sigma} \left(\frac{x - \mu}{\sigma} \right)^{-1-\lambda} \exp \left(- \left(\frac{x - \mu}{\sigma} \right)^{-\lambda} \right)$$

Cumulative distribution function

$$F(x) = \exp \left(- \left(\frac{x - \mu}{\sigma} \right)^{-\lambda} \right)$$

Quantile function

$$F^{-1}(p) = \mu + \sigma - \log(p)^{-1/\lambda}$$

References

Bury, K. (1999). Statistical Distributions in Engineering. Cambridge University Press.

Examples

```
x <- rfrechet(1e5, 5, 2, 1.5)
xx <- seq(0, 1000, by = 0.1)
hist(x, 200, freq = FALSE)
lines(xx, dfrechet(xx, 5, 2, 1.5), col = "red")
hist(pfrechet(x, 5, 2, 1.5))
plot(ecdf(x))
lines(xx, pfrechet(xx, 5, 2, 1.5), col = "red", lwd = 2)
```

GammaPoiss

Gamma-Poisson distribution

Description

Probability mass function and random generation for the gamma-Poisson distribution.

Usage

```
dgpais(x, shape, rate, scale = 1/rate, log = FALSE)
pgpais(q, shape, rate, scale = 1/rate, lower.tail = TRUE, log.p = FALSE)
rgpais(n, shape, rate, scale = 1/rate)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>shape, scale</code>	shape and scale parameters. Must be positive, scale strictly.
<code>rate</code>	an alternative way to specify the scale.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as <code>log(p)</code> .
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Gamma-Poisson distribution arises as a continuous mixture of Poisson distributions, where the mixing distribution of the Poisson rate λ is a gamma distribution. When $X \sim \text{Poisson}(\lambda)$ and $\lambda \sim \text{Gamma}(\alpha, \beta)$, then $X \sim \text{GammaPoisson}(\alpha, \beta)$.

Probability mass function

$$f(x) = \frac{\Gamma(\alpha + x)}{x! \Gamma(\alpha)} \left(\frac{\beta}{1 + \beta} \right)^x \left(1 - \frac{\beta}{1 + \beta} \right)^\alpha$$

Cumulative distribution function is calculated using recursive algorithm that employs the fact that $\Gamma(x) = (x - 1)!$. This enables re-writing probability mass function as

$$f(x) = \frac{(\alpha + x - 1)!}{x! \Gamma(\alpha)} \left(\frac{\beta}{1 + \beta} \right)^x \left(1 - \frac{\beta}{1 + \beta} \right)^\alpha$$

what makes recursive updating from x to $x + 1$ easy using the properties of factorials

$$f(x + 1) = \frac{(\alpha + x - 1)! (\alpha + x)}{x! (x + 1) \Gamma(\alpha)} \left(\frac{\beta}{1 + \beta} \right)^x \left(\frac{\beta}{1 + \beta} \right) \left(1 - \frac{\beta}{1 + \beta} \right)^\alpha$$

and let's us efficiently calculate cumulative distribution function as a sum of probability mass functions

$$F(x) = \sum_{k=0}^x f(k)$$

See Also

[Gamma, Poisson](#)

Examples

```
x <- rgpois(1e5, 7, 0.002)
xx <- seq(0, 12000, by = 1)
hist(x, 100, freq = FALSE)
lines(xx, dgpois(xx, 7, 0.002), col = "red")
hist(pgpois(x, 7, 0.002))
```

```
xx <- seq(0, 12000, by = 0.1)
plot(ecdf(xx))
lines(xx, pppois(xx, 7, 0.002), col = "red", lwd = 2)
```

GEV

*Generalized extreme value distribution***Description**

Density, distribution function, quantile function and random generation for the generalized extreme value distribution.

Usage

```
dgev(x, mu = 0, sigma = 1, xi = 0, log = FALSE)

pgev(q, mu = 0, sigma = 1, xi = 0, lower.tail = TRUE, log.p = FALSE)

qgev(p, mu = 0, sigma = 1, xi = 0, lower.tail = TRUE, log.p = FALSE)

rgev(n, mu = 0, sigma = 1, xi = 0)
```

Arguments

x, q	vector of quantiles.
mu, sigma, xi	location, scale, and shape parameters. Scale must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \frac{1}{\sigma} (1 + \xi \frac{x-\mu}{\sigma})^{-1/\xi-1} \exp\left(- (1 + \xi \frac{x-\mu}{\sigma})^{-1/\xi}\right) & \xi \neq 0 \\ \frac{1}{\sigma} \exp\left(-\frac{x-\mu}{\sigma}\right) \exp\left(-\exp\left(-\frac{x-\mu}{\sigma}\right)\right) & \xi = 0 \end{cases}$$

Cumulative distribution function

$$F(x) = \begin{cases} \exp\left(- (1 + \xi \frac{x-\mu}{\sigma})^{1/\xi}\right) & \xi \neq 0 \\ \exp\left(-\exp\left(-\frac{x-\mu}{\sigma}\right)\right) & \xi = 0 \end{cases}$$

Quantile function

$$F^{-1}(p) = \begin{cases} \mu - \frac{\sigma}{\xi} (1 - (-\log(p))^\xi) & \xi \neq 0 \\ \mu - \sigma \log(-\log(p)) & \xi = 0 \end{cases}$$

References

Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values. Springer.

Examples

```
curve(dgev(x, xi = -1/2), -4, 4, col = "green", ylab = "")
curve(dgev(x, xi = 0), -4, 4, col = "red", add = TRUE)
curve(dgev(x, xi = 1/2), -4, 4, col = "blue", add = TRUE)
legend("topleft", col = c("green", "red", "blue"), lty = 1,
      legend = expression(xi == -1/2, xi == 0, xi == 1/2), bty = "n")

x <- rgev(1e5, 5, 2, .5)
hist(x, 1000, freq = FALSE, xlim = c(0, 50))
curve(dgev(x, 5, 2, .5), 0, 50, col = "red", add = TRUE, n = 5000)
hist(pgev(x, 5, 2, .5))
plot(ecdf(x), xlim = c(0, 50))
curve(pgev(x, 5, 2, .5), 0, 50, col = "red", lwd = 2, add = TRUE)
```

Gompertz

Gompertz distribution

Description

Density, distribution function, quantile function and random generation for the Gompertz distribution.

Usage

```
dgompertz(x, a = 1, b = 1, log = FALSE)

pgompertz(q, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)

qgompertz(p, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)

rgompertz(n, a = 1, b = 1)
```

Arguments

x, q	vector of quantiles.
a, b	positive valued scale and location parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = a \exp\left(bx - \frac{a}{b}(\exp(bx) - 1)\right)$$

Cumulative distribution function

$$F(x) = 1 - \exp\left(-\frac{a}{b}(\exp(bx) - 1)\right)$$

Quantile function

$$F^{-1}(p) = \frac{1}{b} \log\left(1 - \frac{b}{a} \log(1 - p)\right)$$

References

Lenart, A. (2012). The Gompertz distribution and Maximum Likelihood Estimation of its parameters - a revision. MPIDR WORKING PAPER WP 2012-008. <https://www.demogr.mpg.de/papers/working/wp-2012-008.pdf>

Examples

```
x <- rgompertz(1e5, 5, 2)
hist(x, 100, freq = FALSE)
curve(dgompertz(x, 5, 2), 0, 1, col = "red", add = TRUE)
hist(pgompertz(x, 5, 2))
plot(ecdf(x))
curve(pgompertz(x, 5, 2), 0, 1, col = "red", lwd = 2, add = TRUE)
```

GPD

Generalized Pareto distribution

Description

Density, distribution function, quantile function and random generation for the generalized Pareto distribution.

Usage

```
dgpd(x, mu = 0, sigma = 1, xi = 0, log = FALSE)
pgpd(q, mu = 0, sigma = 1, xi = 0, lower.tail = TRUE, log.p = FALSE)
qgpd(p, mu = 0, sigma = 1, xi = 0, lower.tail = TRUE, log.p = FALSE)
rgpd(n, mu = 0, sigma = 1, xi = 0)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>mu, sigma, xi</code>	location, scale, and shape parameters. Scale must be positive.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \frac{1}{\sigma} \left(1 + \xi \frac{x-\mu}{\sigma}\right)^{-(\xi+1)/\xi} & \xi \neq 0 \\ \frac{1}{\sigma} \exp\left(-\frac{x-\mu}{\sigma}\right) & \xi = 0 \end{cases}$$

Cumulative distribution function

$$F(x) = \begin{cases} 1 - \left(1 + \xi \frac{x-\mu}{\sigma}\right)^{-1/\xi} & \xi \neq 0 \\ 1 - \exp\left(-\frac{x-\mu}{\sigma}\right) & \xi = 0 \end{cases}$$

Quantile function

$$F^{-1}(x) = \begin{cases} \mu + \sigma \frac{(1-p)^{-\xi} - 1}{\xi} & \xi \neq 0 \\ \mu - \sigma \log(1-p) & \xi = 0 \end{cases}$$

References

Coles, S. (2001). An Introduction to Statistical Modeling of Extreme Values. Springer.

Examples

```
x <- rgpd(1e5, 5, 2, .1)
hist(x, 100, freq = FALSE, xlim = c(0, 50))
curve(dgpd(x, 5, 2, .1), 0, 50, col = "red", add = TRUE, n = 5000)
hist(pgpd(x, 5, 2, .1))
plot(ecdf(x))
curve(pgpd(x, 5, 2, .1), 0, 50, col = "red", lwd = 2, add = TRUE)
```

Gumbel

*Gumbel distribution***Description**

Density, distribution function, quantile function and random generation for the Gumbel distribution.

Usage

```
dgumbel(x, mu = 0, sigma = 1, log = FALSE)
```

```
pgumbel(q, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)
```

```
qgumbel(p, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rgumbel(n, mu = 0, sigma = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>mu, sigma</code>	location and scale parameters. Scale must be positive.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{1}{\sigma} \exp\left(-\left(\frac{x - \mu}{\sigma} + \exp\left(-\frac{x - \mu}{\sigma}\right)\right)\right)$$

Cumulative distribution function

$$F(x) = \exp\left(-\exp\left(-\frac{x - \mu}{\sigma}\right)\right)$$

Quantile function

$$F^{-1}(p) = \mu - \sigma \log(-\log(p))$$

References

Bury, K. (1999). *Statistical Distributions in Engineering*. Cambridge University Press.

Examples

```
x <- rgumbel(1e5, 5, 2)
hist(x, 100, freq = FALSE)
curve(dgumbel(x, 5, 2), 0, 25, col = "red", add = TRUE)
hist(pgumbel(x, 5, 2))
plot(ecdf(x))
curve(pgumbel(x, 5, 2), 0, 25, col = "red", lwd = 2, add = TRUE)
```

HalfCauchy

Half-Cauchy distribution

Description

Density, distribution function, quantile function and random generation for the half-Cauchy distribution.

Usage

```
dhcauchy(x, sigma = 1, log = FALSE)
phcauchy(q, sigma = 1, lower.tail = TRUE, log.p = FALSE)
qhcauchy(p, sigma = 1, lower.tail = TRUE, log.p = FALSE)
rhcauchy(n, sigma = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>sigma</code>	positive valued scale parameter.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If X follows Cauchy centered at 0 and parametrized by scale σ , then $|X|$ follows half-Cauchy distribution parametrized by scale σ . Half-Cauchy distribution is a special case of half-t distribution with $\nu = 1$ degrees of freedom.

References

Gelman, A. (2006). Prior distributions for variance parameters in hierarchical models (comment on article by Browne and Draper). *Bayesian analysis*, 1(3), 515-534.

Jacob, E. and Jayakumar, K. (2012). On Half-Cauchy Distribution and Process. *International Journal of Statistika and Matematika*, 3(2), 77-81.

See Also

[HalfT](#)

Examples

```
x <- rhcauchy(1e5, 2)
hist(x, 2e5, freq = FALSE, xlim = c(0, 100))
curve(dhcauchy(x, 2), 0, 100, col = "red", add = TRUE)
hist(phcauchy(x, 2))
plot(ecdf(x), xlim = c(0, 100))
curve(phcauchy(x, 2), col = "red", lwd = 2, add = TRUE)
```

HalfNormal

Half-normal distribution

Description

Density, distribution function, quantile function and random generation for the half-normal distribution.

Usage

```
dhnorm(x, sigma = 1, log = FALSE)

phnorm(q, sigma = 1, lower.tail = TRUE, log.p = FALSE)

qhnorm(p, sigma = 1, lower.tail = TRUE, log.p = FALSE)

rhnorm(n, sigma = 1)
```

Arguments

x, q	vector of quantiles.
sigma	positive valued scale parameter.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If X follows normal distribution centered at 0 and parametrized by scale σ , then $|X|$ follows half-normal distribution parametrized by scale σ . Half-t distribution with $\nu = \infty$ degrees of freedom converges to half-normal distribution.

References

Gelman, A. (2006). Prior distributions for variance parameters in hierarchical models (comment on article by Browne and Draper). *Bayesian analysis*, 1(3), 515-534.

Jacob, E. and Jayakumar, K. (2012). On Half-Cauchy Distribution and Process. *International Journal of Statistika and Matematika*, 3(2), 77-81.

See Also

[HalfT](#)

Examples

```
x <- rnorm(1e5, 2)
hist(x, 100, freq = FALSE)
curve(dhnorm(x, 2), 0, 8, col = "red", add = TRUE)
hist(phnorm(x, 2))
plot(ecdf(x))
curve(phnorm(x, 2), 0, 8, col = "red", lwd = 2, add = TRUE)
```

HalfT

Half-t distribution

Description

Density, distribution function, quantile function and random generation for the half-t distribution.

Usage

```
dht(x, nu, sigma = 1, log = FALSE)

pht(q, nu, sigma = 1, lower.tail = TRUE, log.p = FALSE)

qht(p, nu, sigma = 1, lower.tail = TRUE, log.p = FALSE)

rht(n, nu, sigma = 1)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>nu, sigma</code>	positive valued degrees of freedom and scale parameters.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If X follows t distribution parametrized by degrees of freedom ν and scale σ , then $|X|$ follows half-t distribution parametrized by degrees of freedom ν and scale σ .

References

Gelman, A. (2006). Prior distributions for variance parameters in hierarchical models (comment on article by Browne and Draper). *Bayesian analysis*, 1(3), 515-534.

Jacob, E. and Jayakumar, K. (2012). On Half-Cauchy Distribution and Process. *International Journal of Statistika and Matematika*, 3(2), 77-81.

See Also

[HalfNormal](#), [HalfCauchy](#)

Examples

```
x <- rht(1e5, 2, 2)
hist(x, 500, freq = FALSE, xlim = c(0, 100))
curve(dht(x, 2, 2), 0, 100, col = "red", add = TRUE)
hist(pht(x, 2, 2))
plot(ecdf(x), xlim = c(0, 100))
curve(pht(x, 2, 2), 0, 100, col = "red", lwd = 2, add = TRUE)
```

Huber

"Huber density" distribution

Description

Density, distribution function, quantile function and random generation for the "Huber density" distribution.

Usage

```

dhuber(x, mu = 0, sigma = 1, epsilon = 1.345, log = FALSE)

phuber(q, mu = 0, sigma = 1, epsilon = 1.345, lower.tail = TRUE, log.p = FALSE)

qhuber(p, mu = 0, sigma = 1, epsilon = 1.345, lower.tail = TRUE, log.p = FALSE)

rhuber(n, mu = 0, sigma = 1, epsilon = 1.345)

```

Arguments

x, q	vector of quantiles.
mu, sigma, epsilon	location, and scale, and shape parameters. Scale and shape must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Huber density is connected to Huber loss and can be defined as:

$$f(x) = \frac{1}{2\sqrt{2\pi} \left(\Phi(k) + \phi(k)/k - \frac{1}{2} \right)} e^{-\rho_k(x)}$$

where

$$\rho_k(x) = \begin{cases} \frac{1}{2}x^2 & |x| \leq k \\ k|x| - \frac{1}{2}k^2 & |x| > k \end{cases}$$

References

Huber, P.J. (1964). Robust Estimation of a Location Parameter. *Annals of Statistics*, 53(1), 73-101.

Huber, P.J. (1981). *Robust Statistics*. Wiley.

Schumann, D. (2009). *Robust Variable Selection*. ProQuest.

Examples

```

x <- rhuber(1e5, 5, 2, 3)
hist(x, 100, freq = FALSE)
curve(dhuber(x, 5, 2, 3), -20, 20, col = "red", add = TRUE, n = 5000)
hist(phuber(x, 5, 2, 3))
plot(ecdf(x))
curve(phuber(x, 5, 2, 3), -20, 20, col = "red", lwd = 2, add = TRUE)

```

 InvChiSq

Inverse chi-squared and scaled chi-squared distributions

Description

Density, distribution function and random generation for the inverse chi-squared distribution and scaled chi-squared distribution. The implementation uses a non-standard linear scale parameterization of τ ; see Details for the precise definition and its relation to the standard τ^2 formulation.

Usage

```
dinvchisq(x, nu, tau, log = FALSE)
```

```
pinvchisq(q, nu, tau, lower.tail = TRUE, log.p = FALSE)
```

```
qinvchisq(p, nu, tau, lower.tail = TRUE, log.p = FALSE)
```

```
rinvchisq(n, nu, tau)
```

Arguments

x, q	vector of quantiles.
nu	positive valued shape parameter.
tau	positive valued scaling parameter; if provided it returns values for scaled chi-squared distributions.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If X follows $\chi^2(\nu)$ distribution, then $1/X$ follows inverse chi-squared distribution parametrized by ν . Inverse chi-squared distribution is a special case of inverse gamma distribution with parameters $\alpha = \frac{\nu}{2}$ and $\beta = \frac{1}{2}$; or $\alpha = \frac{\nu}{2}$ and $\beta = \frac{\nu\tau}{2}$.

Note on parameterization: this implementation uses a linear scaling parameter τ . The scaled inverse chi-squared distribution is defined via the equivalent inverse-gamma form

$$X \sim \text{InvGamma}\left(\frac{\nu}{2}, \frac{\nu\tau}{2}\right),$$

where τ enters linearly in the second parameter.

This differs from the most common statistical convention, which uses a squared scale parameter τ^2 :

$$X \sim \text{InvGamma}\left(\frac{\nu}{2}, \frac{\nu\tau^2}{2}\right),$$

The two parameterizations are related by

$$\tau_{\text{thisimplementation}} = \tau_{\text{standard}}^2.$$

Users familiar with the standard scaled inverse chi-squared distribution should be aware that τ here corresponds to a variance-scale parameter rather than a standard-deviation-type parameter.

See Also

[Chisquare](#), [GammaDist](#)

Examples

```
x <- rinvchisq(1e5, 20)
hist(x, 100, freq = FALSE)
curve(dinvchisq(x, 20), 0, 1, n = 501, col = "red", add = TRUE)
hist(pinvchisq(x, 20))
plot(ecdf(x))
curve(pinvchisq(x, 20), 0, 1, n = 501, col = "red", lwd = 2, add = TRUE)

# scaled

x <- rinvchisq(1e5, 10, 5)
hist(x, 100, freq = FALSE)
curve(dinvchisq(x, 10, 5), 0, 150, n = 501, col = "red", add = TRUE)
hist(pinvchisq(x, 10, 5))
plot(ecdf(x))
curve(pinvchisq(x, 10, 5), 0, 150, n = 501, col = "red", lwd = 2, add = TRUE)
```

InvGamma

Inverse-gamma distribution

Description

Density, distribution function and random generation for the inverse-gamma distribution.

Usage

```
dinvgamma(x, alpha, beta = 1, log = FALSE)

pinvgamma(q, alpha, beta = 1, lower.tail = TRUE, log.p = FALSE)

qinvgamma(p, alpha, beta = 1, lower.tail = TRUE, log.p = FALSE)

rinvgamma(n, alpha, beta = 1)
```

Arguments

x, q	vector of quantiles.
alpha, beta	positive valued shape and scale parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability mass function

$$f(x) = \frac{\beta^\alpha x^{-\alpha-1} \exp(-\frac{\beta}{x})}{\Gamma(\alpha)}$$

Cumulative distribution function

$$F(x) = \frac{\gamma(\alpha, \frac{\beta}{x})}{\Gamma(\alpha)}$$

References

Witkovsky, V. (2001). Computing the distribution of a linear combination of inverted gamma variables. *Kybernetika* 37(1), 79-90.

Leemis, L.M. and McQueston, L.T. (2008). Univariate Distribution Relationships. *American Statistician* 62(1): 45-53.

See Also

[GammaDist](#)

Examples

```
x <- rinvgamma(1e5, 20, 3)
hist(x, 100, freq = FALSE)
curve(dinvgamma(x, 20, 3), 0, 1, col = "red", add = TRUE, n = 5000)
hist(pinvgamma(x, 20, 3))
plot(ecdf(x))
curve(pinvgamma(x, 20, 3), 0, 1, col = "red", lwd = 2, add = TRUE, n = 5000)
```

Kumaraswamy

*Kumaraswamy distribution***Description**

Density, distribution function, quantile function and random generation for the Kumaraswamy distribution.

Usage

```
dkumar(x, a = 1, b = 1, log = FALSE)
```

```
pkumar(q, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)
```

```
qkumar(p, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rkumar(n, a = 1, b = 1)
```

Arguments

x, q	vector of quantiles.
a, b	positive valued parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = abx^{a-1}(1-x^a)^{b-1}$$

Cumulative distribution function

$$F(x) = 1 - (1-x^a)^b$$

Quantile function

$$F^{-1}(p) = 1 - (1-p^{1/b})^{1/a}$$

References

- Jones, M. C. (2009). Kumaraswamy's distribution: A beta-type distribution with some tractability advantages. *Statistical Methodology*, 6, 70-81.
- Cordeiro, G.M. and de Castro, M. (2009). A new family of generalized distributions. *Journal of Statistical Computation & Simulation*, 1-17.

Examples

```
x <- rkumar(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dkumar(x, 5, 16), 0, 1, col = "red", add = TRUE)
hist(pkumar(x, 5, 16))
plot(ecdf(x))
curve(pkumar(x, 5, 16), 0, 1, col = "red", lwd = 2, add = TRUE)
```

Laplace

*Laplace distribution***Description**

Density, distribution function, quantile function and random generation for the Laplace distribution.

Usage

```
dlaplace(x, mu = 0, sigma = 1, log = FALSE)

plaplace(q, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)

qlaplace(p, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)

rlaplace(n, mu = 0, sigma = 1)
```

Arguments

x, q	vector of quantiles.
mu, sigma	location and scale parameters. Scale must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{1}{2\sigma} \exp\left(-\left|\frac{x-\mu}{\sigma}\right|\right)$$

Cumulative distribution function

$$F(x) = \begin{cases} \frac{1}{2} \exp\left(\frac{x-\mu}{\sigma}\right) & x < \mu \\ 1 - \frac{1}{2} \exp\left(\frac{x-\mu}{\sigma}\right) & x \geq \mu \end{cases}$$

Quantile function

$$F^{-1}(p) = \begin{cases} \mu + \sigma \log(2p) & p < 0.5 \\ \mu - \sigma \log(2(1-p)) & p \geq 0.5 \end{cases}$$

References

- Krishnamoorthy, K. (2006). Handbook of Statistical Distributions with Applications. Chapman & Hall/CRC
- Forbes, C., Evans, M. Hastings, N., & Peacock, B. (2011). Statistical Distributions. John Wiley & Sons.

Examples

```
x <- rlaplace(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dlaplace(x, 5, 16), -200, 200, n = 500, col = "red", add = TRUE)
hist(plaplace(x, 5, 16))
plot(ecdf(x))
curve(plaplace(x, 5, 16), -200, 200, n = 500, col = "red", lwd = 2, add = TRUE)
```

LocationScaleT

Location-scale version of the t-distribution

Description

Probability mass function, distribution function and random generation for location-scale version of the t-distribution. Location-scale version of the t-distribution besides degrees of freedom ν , is parametrized using additional parameters μ for location and σ for scale ($\mu = 0$ and $\sigma = 1$ for standard t-distribution).

Usage

```
dlst(x, df, mu = 0, sigma = 1, log = FALSE)

plst(q, df, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)

qlst(p, df, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)

rlst(n, df, mu = 0, sigma = 1)
```

Arguments

x, q	vector of quantiles.
df	degrees of freedom (> 0 , maybe non-integer). df = Inf is allowed.
mu	vector of locations
sigma	vector of positive valued scale parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If length(n) > 1 , the length is taken to be the number required.

See Also[TDist](#)**Examples**

```
x <- rlst(1e5, 1000, 5, 13)
hist(x, 100, freq = FALSE)
curve(dlst(x, 1000, 5, 13), -60, 60, col = "red", add = TRUE)
hist(plst(x, 1000, 5, 13))
plot(ecdf(x))
curve(plst(x, 1000, 5, 13), -60, 60, col = "red", lwd = 2, add = TRUE)
```

LogSeries

*Logarithmic series distribution***Description**

Density, distribution function, quantile function and random generation for the logarithmic series distribution.

Usage

```
dlgser(x, theta, log = FALSE)

plgser(q, theta, lower.tail = TRUE, log.p = FALSE)

qlgser(p, theta, lower.tail = TRUE, log.p = FALSE)

rlgser(n, theta)
```

Arguments

x, q	vector of quantiles.
theta	vector; concentration parameter; ($0 < \text{theta} < 1$).
log, log.p	logical; if TRUE, probabilities p are given as $\log(p)$.
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability mass function

$$f(x) = \frac{-1}{\log(1 - \theta)} \frac{\theta^x}{x}$$

Cumulative distribution function

$$F(x) = \frac{-1}{\log(1 - \theta)} \sum_{k=1}^x \frac{\theta^k}{k}$$

Quantile function and random generation are computed using algorithm described in Krishnamoorthy (2006).

References

Krishnamoorthy, K. (2006). Handbook of Statistical Distributions with Applications. Chapman & Hall/CRC

Forbes, C., Evans, M. Hastings, N., & Peacock, B. (2011). Statistical Distributions. John Wiley & Sons.

Examples

```
x <- rlgser(1e5, 0.66)
xx <- seq(0, 100, by = 1)
plot(prop.table(table(x)), type = "h")
lines(xx, dlgser(xx, 0.66), col = "red")

# Notice: distribution of F(X) is far from uniform:
hist(plgser(x, 0.66), 50)

xx <- seq(0, 100, by = 0.01)
plot(ecdf(x))
lines(xx, plgser(xx, 0.66), col = "red", lwd = 2)
```

Lomax

Lomax distribution

Description

Density, distribution function, quantile function and random generation for the Lomax distribution.

Usage

```
dlomax(x, lambda, kappa, log = FALSE)
plomax(q, lambda, kappa, lower.tail = TRUE, log.p = FALSE)
qlomax(p, lambda, kappa, lower.tail = TRUE, log.p = FALSE)
rlomax(n, lambda, kappa)
```

Arguments

`x, q` vector of quantiles.

`lambda, kappa` positive valued parameters.

`log, log.p` logical; if TRUE, probabilities `p` are given as $\log(p)$.

`lower.tail` logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.

`p` vector of probabilities.

`n` number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{\lambda\kappa}{(1 + \lambda x)^{\kappa+1}}$$

Cumulative distribution function

$$F(x) = 1 - (1 + \lambda x)^{-\kappa}$$

Quantile function

$$F^{-1}(p) = \frac{(1 - p)^{-1/\kappa} - 1}{\lambda}$$

Examples

```
x <- rlomax(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dlomax(x, 5, 16), 0, 1, col = "red", add = TRUE, n = 5000)
hist(plomax(x, 5, 16))
plot(ecdf(x))
curve(plomax(x, 5, 16), 0, 1, col = "red", lwd = 2, add = TRUE)
```

MultiHypergeometric *Multivariate hypergeometric distribution*

Description

Probability mass function and random generation for the multivariate hypergeometric distribution.

Usage

```
dmvhyper(x, n, k, log = FALSE)
```

```
rmvhyper(nn, n, k)
```

Arguments

x	<i>m</i> -column matrix of quantiles.
n	<i>m</i> -length vector or <i>m</i> -column matrix of numbers of balls in <i>m</i> colors.
k	the number of balls drawn from the urn.
log	logical; if TRUE, probabilities p are given as log(p).
nn	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

Probability mass function

$$f(x) = \frac{\prod_{i=1}^m \binom{n_i}{x_i}}{\binom{N}{k}}$$

The multivariate hypergeometric distribution is generalization of hypergeometric distribution. It is used for sampling *without* replacement *k* out of *N* marbles in *m* colors, where each of the colors appears n_i times. Where $k = \sum_{i=1}^m x_i$, $N = \sum_{i=1}^m n_i$ and $k \leq N$.

References

Gentle, J.E. (2006). Random number generation and Monte Carlo methods. Springer.

See Also

[Hypergeometric](#)

Examples

```
# Generating 10 random draws from multivariate hypergeometric
# distribution parametrized using a vector

rmvhyper(10, c(10, 12, 5, 8, 11), 33)
```

 Multinomial

Multinomial distribution

Description

Probability mass function and random generation for the multinomial distribution.

Usage

```
dmnom(x, size, prob, log = FALSE)
```

```
rmnom(n, size, prob)
```

Arguments

x	<i>k</i> -column matrix of quantiles.
size	numeric vector; number of trials (zero or more).
prob	<i>k</i> -column numeric matrix; probability of success on each trial.
log	logical; if TRUE, probabilities p are given as log(p).
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

Probability mass function

$$f(x) = \frac{n!}{\prod_{i=1}^k x_i} \prod_{i=1}^k p_i^{x_i}$$

References

Gentle, J.E. (2006). Random number generation and Monte Carlo methods. Springer.

See Also

[Binomial](#), [Multinomial](#)

Examples

```
# Generating 10 random draws from multinomial distribution
# parametrized using a vector

(x <- rmnom(10, 3, c(1/3, 1/3, 1/3)))

# Results are consistent with dmultinom() from stats:

all.equal(dmultinom(x[1,], 3, c(1/3, 1/3, 1/3)),
          dmnom(x[1, , drop = FALSE], 3, c(1/3, 1/3, 1/3)))
```

NegHyper

*Negative hypergeometric distribution***Description**

Probability mass function, distribution function, quantile function and random generation for the negative hypergeometric distribution.

Usage

```
dnhyper(x, n, m, r, log = FALSE)
```

```
pnhyper(q, n, m, r, lower.tail = TRUE, log.p = FALSE)
```

```
qnhyper(p, n, m, r, lower.tail = TRUE, log.p = FALSE)
```

```
rnhyper(nn, n, m, r)
```

Arguments

x, q	vector of quantiles representing the number of balls drawn without replacement from an urn which contains both black and white balls.
n	the number of black balls in the urn.
m	the number of white balls in the urn.
r	the number of white balls that needs to be drawn for the sampling to be stopped.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
nn	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Negative hypergeometric distribution describes number of balls x observed until drawing without replacement to obtain r white balls from the urn containing m white balls and n black balls, and is defined as

$$f(x) = \frac{\binom{x-1}{r-1} \binom{m+n-x}{m-r}}{\binom{m+n}{n}}$$

The algorithm used for calculating probability mass function, cumulative distribution function and quantile function is based on Fortran program NHYPERG created by Berry and Mielke (1996, 1998). Random generation is done by inverse transform sampling.

References

- Berry, K. J., & Mielke, P. W. (1998). The negative hypergeometric probability distribution: Sampling without replacement from a finite population. *Perceptual and motor skills*, 86(1), 207-210. <https://journals.sagepub.com/doi/10.2466/pms.1998.86.1.207>
- Berry, K. J., & Mielke, P. W. (1996). Exact confidence limits for population proportions based on the negative hypergeometric probability distribution. *Perceptual and motor skills*, 83(3 suppl), 1216-1218. <https://journals.sagepub.com/doi/10.2466/pms.1996.83.3f.1216>
- Schuster, E. F., & Sype, W. R. (1987). On the negative hypergeometric distribution. *International Journal of Mathematical Education in Science and Technology*, 18(3), 453-459.
- Chae, K. C. (1993). Presenting the negative hypergeometric distribution to the introductory statistics courses. *International Journal of Mathematical Education in Science and Technology*, 24(4), 523-526.
- Jones, S.N. (2013). A Gaming Application of the Negative Hypergeometric Distribution. UNLV Theses, Dissertations, Professional Papers, and Capstones. Paper 1846. <https://oasis.library.unlv.edu/cgi/viewcontent.cgi?referer=&httpsredir=1&article=2847&context=thesesdissertations>

See Also

[Hypergeometric](#)

Examples

```
x <- rnhyper(1e5, 60, 35, 15)
xx <- 15:95
plot(prop.table(table(x)))
lines(xx, dhyper(xx, 60, 35, 15), col = "red")
hist(pnhyper(x, 60, 35, 15))

xx <- seq(0, 100, by = 0.01)
plot(ecdf(x))
lines(xx, pnhyper(xx, 60, 35, 15), col = "red", lwd = 2)
```

NormalMix

Mixture of normal distributions

Description

Density, distribution function and random generation for the mixture of normal distributions.

Usage

```
dmixnorm(x, mean, sd, alpha, log = FALSE)

pmixnorm(q, mean, sd, alpha, lower.tail = TRUE, log.p = FALSE)

rmixnorm(n, mean, sd, alpha)
```

Arguments

x, q	vector of quantiles.
mean	matrix (or vector) of means.
sd	matrix (or vector) of standard deviations.
alpha	matrix (or vector) of mixing proportions; mixing proportions need to sum up to 1.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.
p	vector of probabilities.

Details

Probability density function

$$f(x) = \alpha_1 f_1(x; \mu_1, \sigma_1) + \dots + \alpha_k f_k(x; \mu_k, \sigma_k)$$

Cumulative distribution function

$$F(x) = \alpha_1 F_1(x; \mu_1, \sigma_1) + \dots + \alpha_k F_k(x; \mu_k, \sigma_k)$$

where $\sum_i \alpha_i = 1$.

Examples

```
x <- rmixnorm(1e5, c(0.5, 3, 6), c(3, 1, 1), c(1/3, 1/3, 1/3))
hist(x, 100, freq = FALSE)
curve(dmixnorm(x, c(0.5, 3, 6), c(3, 1, 1), c(1/3, 1/3, 1/3)),
      -20, 20, n = 500, col = "red", add = TRUE)
hist(pmixnorm(x, c(0.5, 3, 6), c(3, 1, 1), c(1/3, 1/3, 1/3)))
plot(ecdf(x))
curve(pmixnorm(x, c(0.5, 3, 6), c(3, 1, 1), c(1/3, 1/3, 1/3)),
      -20, 20, n = 500, col = "red", lwd = 2, add = TRUE)
```

Description

Non-standard form of beta distribution with lower and upper bounds denoted as min and max. By default min=0 and max=1 what leads to standard beta distribution.

Usage

```
dnsbeta(x, shape1, shape2, min = 0, max = 1, log = FALSE)
pnsbeta(q, shape1, shape2, min = 0, max = 1, lower.tail = TRUE, log.p = FALSE)
qnsbeta(p, shape1, shape2, min = 0, max = 1, lower.tail = TRUE, log.p = FALSE)
rnsbeta(n, shape1, shape2, min = 0, max = 1)
```

Arguments

x, q	vector of quantiles.
shape1, shape2	non-negative parameters of the Beta distribution.
min, max	lower and upper bounds.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$, otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

See Also

[Beta](#)

Examples

```
x <- rnsbeta(1e5, 5, 13, -4, 8)
hist(x, 100, freq = FALSE)
curve(dnsbeta(x, 5, 13, -4, 8), -4, 6, col = "red", add = TRUE)
hist(pnsbeta(x, 5, 13, -4, 8))
plot(ecdf(x))
curve(pnsbeta(x, 5, 13, -4, 8), -4, 6, col = "red", lwd = 2, add = TRUE)
```

Pareto

Pareto distribution

Description

Density, distribution function, quantile function and random generation for the Pareto distribution.

Usage

```
dpareto(x, a = 1, b = 1, log = FALSE)
ppareto(q, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)
qpareto(p, a = 1, b = 1, lower.tail = TRUE, log.p = FALSE)
rpareto(n, a = 1, b = 1)
```

Arguments

`x, q` vector of quantiles.
`a, b` positive valued scale and location parameters.
`log, log.p` logical; if TRUE, probabilities `p` are given as $\log(p)$.
`lower.tail` logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
`p` vector of probabilities.
`n` number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{ab^a}{x^{a+1}}$$

Cumulative distribution function

$$F(x) = 1 - \left(\frac{b}{x}\right)^a$$

Quantile function

$$F^{-1}(p) = \frac{b}{(1-p)^{1-a}}$$

References

Krishnamoorthy, K. (2006). Handbook of Statistical Distributions with Applications. Chapman & Hall/CRC

Examples

```
x <- rpareto(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dpareto(x, 5, 16), 0, 200, col = "red", add = TRUE)
hist(ppareto(x, 5, 16))
plot(ecdf(x))
curve(ppareto(x, 5, 16), 0, 200, col = "red", lwd = 2, add = TRUE)
```

PoissonMix

*Mixture of Poisson distributions***Description**

Density, distribution function and random generation for the mixture of Poisson distributions.

Usage

```
dmixpois(x, lambda, alpha, log = FALSE)
```

```
pmixpois(q, lambda, alpha, lower.tail = TRUE, log.p = FALSE)
```

```
rmixpois(n, lambda, alpha)
```

Arguments

x, q	vector of quantiles.
lambda	matrix (or vector) of (non-negative) means.
alpha	matrix (or vector) of mixing proportions; mixing proportions need to sum up to 1.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.
p	vector of probabilities.

Details

Probability density function

$$f(x) = \alpha_1 f_1(x; \lambda_1) + \cdots + \alpha_k f_k(x; \lambda_k)$$

Cumulative distribution function

$$F(x) = \alpha_1 F_1(x; \lambda_1) + \cdots + \alpha_k F_k(x; \lambda_k)$$

where $\sum_i \alpha_i = 1$.

Examples

```
x <- rmixpois(1e5, c(5, 12, 19), c(1/3, 1/3, 1/3))
xx <- seq(-1, 50)
plot(prop.table(table(x)))
lines(xx, dmixpois(xx, c(5, 12, 19), c(1/3, 1/3, 1/3)), col = "red")
hist(pmixpois(x, c(5, 12, 19), c(1/3, 1/3, 1/3)))
```

```
xx <- seq(0, 50, by = 0.01)
plot(ecdf(xx))
lines(xx, pmixpois(xx, c(5, 12, 19), c(1/3, 1/3, 1/3)), col = "red", lwd = 2)
```

PowerDist

*Power distribution***Description**

Density, distribution function, quantile function and random generation for the power distribution.

Usage

```
dpower(x, alpha, beta, log = FALSE)
ppower(q, alpha, beta, lower.tail = TRUE, log.p = FALSE)
qpower(p, alpha, beta, lower.tail = TRUE, log.p = FALSE)
rpower(n, alpha, beta)
```

Arguments

x, q	vector of quantiles.
alpha, beta	parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{\beta x^{\beta-1}}{\alpha^\beta}$$

Cumulative distribution function

$$F(x) = \frac{x^\beta}{\alpha^\beta}$$

Quantile function

$$F^{-1}(p) = \alpha p^{1/\beta}$$

Examples

```
x <- rpower(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dpower(x, 5, 16), 2, 6, col = "red", add = TRUE, n = 5000)
hist(ppower(x, 5, 16))
plot(ecdf(x))
curve(ppower(x, 5, 16), 2, 6, col = "red", lwd = 2, add = TRUE)
```

 PropBeta

Beta distribution of proportions

Description

Probability mass function, distribution function and random generation for the reparametrized beta distribution.

Usage

```
dprop(x, size, mean, prior = 0, log = FALSE)
pprop(q, size, mean, prior = 0, lower.tail = TRUE, log.p = FALSE)
qprop(p, size, mean, prior = 0, lower.tail = TRUE, log.p = FALSE)
rprop(n, size, mean, prior = 0)
```

Arguments

x, q	vector of quantiles.
size	non-negative real number; precision or number of binomial trials.
mean	mean proportion or probability of success on each trial; $0 < \text{mean} < 1$.
prior	(see below) with prior = 0 (default) the distribution corresponds to re-parametrized beta distribution used in beta regression. This parameter needs to be non-negative.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Beta can be understood as a distribution of $x = k/\phi$ proportions in ϕ trials where the average proportion is denoted as μ , so its parameters become $\alpha = \phi\mu$ and $\beta = \phi(1 - \mu)$ and its density function becomes

$$f(x) = \frac{x^{\phi\mu + \pi - 1}(1 - x)^{\phi(1 - \mu) + \pi - 1}}{B(\phi\mu + \pi, \phi(1 - \mu) + \pi)}$$

where π is a *prior* parameter, so the distribution is a *posterior* distribution after observing $\phi\mu$ successes and $\phi(1 - \mu)$ failures in ϕ trials with binomial likelihood and symmetric $\text{Beta}(\pi, \pi)$ prior for probability of success. Parameter value $\pi = 1$ corresponds to uniform prior; $\pi = 1/2$ corresponds to Jeffreys prior; $\pi = 0$ corresponds to "uninformative" Haldane prior, this is also the re-parametrized distribution used in beta regression. With $\pi = 0$ the distribution can be understood as a continuous analog to binomial distribution dealing with proportions rather than counts. Alternatively ϕ may be understood as precision parameter (as in beta regression).

Notice that in pre-1.8.4 versions of this package, prior was not settable and by default fixed to one, instead of zero. To obtain the same results as in the previous versions, use `prior = 1` in each of the functions.

References

Ferrari, S., & Cribari-Neto, F. (2004). Beta regression for modelling rates and proportions. *Journal of Applied Statistics*, 31(7), 799-815.

Smithson, M., & Verkuilen, J. (2006). A better lemon squeezer? Maximum-likelihood regression with beta-distributed dependent variables. *Psychological Methods*, 11(1), 54-71.

See Also

[beta](#), [binomial](#)

Examples

```
x <- rprop(1e5, 100, 0.33)
hist(x, 100, freq = FALSE)
curve(dprop(x, 100, 0.33), 0, 1, col = "red", add = TRUE)
hist(pprop(x, 100, 0.33))
plot(ecdf(x))
curve(pprop(x, 100, 0.33), 0, 1, col = "red", lwd = 2, add = TRUE)

n <- 500
p <- 0.23
k <- rbinom(1e5, n, p)
hist(k/n, freq = FALSE, 100)
curve(dprop(x, n, p), 0, 1, col = "red", add = TRUE, n = 500)
```

Rademacher	<i>Random generation from Rademacher distribution</i>
------------	-------------------------------------------------------

Description

Random generation for the Rademacher distribution (values -1 and +1 with equal probability).

Usage

```
rsign(n)
```

Arguments

n	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.
---	-----------------------------------------------------------------------------------------------------------

Rayleigh	<i>Rayleigh distribution</i>
----------	------------------------------

Description

Density, distribution function, quantile function and random generation for the Rayleigh distribution.

Usage

```
drayleigh(x, sigma = 1, log = FALSE)
```

```
prayleigh(q, sigma = 1, lower.tail = TRUE, log.p = FALSE)
```

```
qrayleigh(p, sigma = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rrayleigh(n, sigma = 1)
```

Arguments

x, q	vector of quantiles.
sigma	positive valued parameter.
log, log.p	logical; if TRUE, probabilities p are given as $\log(p)$.
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{x}{\sigma^2} \exp\left(-\frac{x^2}{2\sigma^2}\right)$$

Cumulative distribution function

$$F(x) = 1 - \exp\left(-\frac{x^2}{2\sigma^2}\right)$$

Quantile function

$$F^{-1}(p) = \sqrt{-2\sigma^2 \log(1-p)}$$

References

- Krishnamoorthy, K. (2006). Handbook of Statistical Distributions with Applications. Chapman & Hall/CRC.
- Forbes, C., Evans, M. Hastings, N., & Peacock, B. (2011). Statistical Distributions. John Wiley & Sons.

Examples

```
x <- rrayleigh(1e5, 13)
hist(x, 100, freq = FALSE)
curve(drayleigh(x, 13), 0, 60, col = "red", add = TRUE)
hist(prayleigh(x, 13))
plot(ecdf(x))
curve(prayleigh(x, 13), 0, 60, col = "red", lwd = 2, add = TRUE)
```

ShiftGomp

Shifted Gompertz distribution

Description

Density, distribution function, and random generation for the shifted Gompertz distribution.

Usage

```
dsgomp(x, b, eta, log = FALSE)
psgomp(q, b, eta, lower.tail = TRUE, log.p = FALSE)
rsgomp(n, b, eta)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>b, eta</code>	positive valued scale and shape parameters; both need to be positive.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

If X follows exponential distribution parametrized by scale b and Y follows reparametrized Gumbel distribution with cumulative distribution function $F(x) = \exp(-\eta e^{-bx})$ parametrized by scale b and shape η , then $\max(X, Y)$ follows shifted Gompertz distribution parametrized by scale $b > 0$ and shape $\eta > 0$. The above relation is used by `rsgomp` function for random generation from shifted Gompertz distribution.

Probability density function

$$f(x) = be^{-bx} \exp(-\eta e^{-bx}) [1 + \eta(1 - e^{-bx})]$$

Cumulative distribution function

$$F(x) = (1 - e^{-bx}) \exp(-\eta e^{-bx})$$

References

- Bemmaor, A.C. (1994). Modeling the Diffusion of New Durable Goods: Word-of-Mouth Effect Versus Consumer Heterogeneity. [In:] G. Laurent, G.L. Lilien & B. Pras. Research Traditions in Marketing. Boston: Kluwer Academic Publishers. pp. 201-223.
- Jimenez, T.F. and Jodra, P. (2009). A Note on the Moments and Computer Generation of the Shifted Gompertz Distribution. Communications in Statistics - Theory and Methods, 38(1), 78-89.
- Jimenez T.F. (2014). Estimation of the Parameters of the Shifted Gompertz Distribution, Using Least Squares, Maximum Likelihood and Moments Methods. Journal of Computational and Applied Mathematics, 255(1), 867-877.

Examples

```
x <- rsgomp(1e5, 0.4, 1)
hist(x, 50, freq = FALSE)
curve(dsgomp(x, 0.4, 1), 0, 30, col = "red", add = TRUE)
hist(psgomp(x, 0.4, 1))
plot(ecdf(x))
curve(psgomp(x, 0.4, 1), 0, 30, col = "red", lwd = 2, add = TRUE)
```

 Skellam

Skellam distribution

Description

Probability mass function and random generation for the Skellam distribution.

Usage

```
dskellam(x, mu1, mu2, log = FALSE)
```

```
rskellam(n, mu1, mu2)
```

Arguments

x	vector of quantiles.
mu1, mu2	positive valued parameters.
log	logical; if TRUE, probabilities p are given as log(p).
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

If X and Y follow Poisson distributions with means μ_1 and μ_2 , then $X - Y$ follows Skellam distribution parametrized by μ_1 and μ_2 .

Probability mass function

$$f(x) = e^{-(\mu_1 + \mu_2)} \left(\frac{\mu_1}{\mu_2} \right)^{k/2} I_k(2\sqrt{\mu_1 \mu_2})$$

References

Karlis, D., & Ntzoufras, I. (2006). Bayesian analysis of the differences of count data. *Statistics in medicine*, 25(11), 1885-1905.

Skellam, J.G. (1946). The frequency distribution of the difference between two Poisson variates belonging to different populations. *Journal of the Royal Statistical Society, series A*, 109(3), 26.

Examples

```
x <- rskellam(1e5, 5, 13)
xx <- -40:40
plot(prop.table(table(x)), type = "h")
lines(xx, dskellam(xx, 5, 13), col = "red")
```

Slash

*Slash distribution***Description**

Probability mass function, distribution function and random generation for slash distribution.

Usage

```
dslash(x, mu = 0, sigma = 1, log = FALSE)
```

```
pslash(q, mu = 0, sigma = 1, lower.tail = TRUE, log.p = FALSE)
```

```
rslash(n, mu = 0, sigma = 1)
```

Arguments

x, q	vector of quantiles.
mu	vector of locations
sigma	vector of positive valued scale parameters.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If length(n) > 1, the length is taken to be the number required.

Details

If $Z \sim \text{Normal}(0, 1)$ and $U \sim \text{Uniform}(0, 1)$, then Z/U follows slash distribution.

Probability density function

$$f(x) = \begin{cases} \frac{\phi(0) - \phi(x)}{x^2} & x \neq 0 \\ \frac{1}{2\sqrt{2\pi}} & x = 0 \end{cases}$$

Cumulative distribution function

$$F(x) = \begin{cases} \Phi(x) - \frac{\phi(0) - \phi(x)}{x} & x \neq 0 \\ \frac{1}{2} & x = 0 \end{cases}$$

Examples

```
x <- rslash(1e5, 5, 3)
hist(x, 1e5, freq = FALSE, xlim = c(-100, 100))
curve(dslash(x, 5, 3), -100, 100, col = "red", n = 500, add = TRUE)
hist(pslash(x, 5, 3))
plot(ecdf(x), xlim = c(-100, 100))
curve(pslash(x, 5, 3), -100, 100, col = "red", lwd = 2, n = 500, add = TRUE)
```

Triangular	<i>Triangular distribution</i>
------------	--------------------------------

Description

Density, distribution function, quantile function and random generation for the triangular distribution.

Usage

`dtriang(x, a = -1, b = 1, c = (a + b)/2, log = FALSE)`

`ptriang(q, a = -1, b = 1, c = (a + b)/2, lower.tail = TRUE, log.p = FALSE)`

`qtriang(p, a = -1, b = 1, c = (a + b)/2, lower.tail = TRUE, log.p = FALSE)`

`rtriang(n, a = -1, b = 1, c = (a + b)/2)`

Arguments

<code>x, q</code>	vector of quantiles.
<code>a, b, c</code>	minimum, maximum and mode of the distribution.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \frac{2(x-a)}{(b-a)(c-a)} & x < c \\ \frac{2}{b-a} & x = c \\ \frac{2(b-x)}{(b-a)(b-c)} & x > c \end{cases}$$

Cumulative distribution function

$$F(x) = \begin{cases} \frac{(x-a)^2}{(b-a)(c-a)} & x \leq c \\ 1 - \frac{(b-x)^2}{(b-a)(b-c)} & x > c \end{cases}$$

Quantile function

$$F^{-1}(p) = \begin{cases} a + \sqrt{p \times (b-a)(c-a)} & p \leq \frac{c-a}{b-a} \\ b - \sqrt{(1-p)(b-a)(b-c)} & p > \frac{c-a}{b-a} \end{cases}$$

For random generation MINMAX method described by Stein and Keblis (2009) is used.

References

- Forbes, C., Evans, M. Hastings, N., & Peacock, B. (2011). *Statistical Distributions*. John Wiley & Sons.
- Stein, W. E., & Kebelis, M. F. (2009). A new method to simulate the triangular distribution. *Mathematical and computer modelling*, 49(5), 1143-1147.

Examples

```
x <- rtriang(1e5, 5, 7, 6)
hist(x, 100, freq = FALSE)
curve(dtriang(x, 5, 7, 6), 3, 10, n = 500, col = "red", add = TRUE)
hist(ptriang(x, 5, 7, 6))
plot(ecdf(x))
curve(ptriang(x, 5, 7, 6), 3, 10, n = 500, col = "red", lwd = 2, add = TRUE)
```

TruncBinom

Truncated binomial distribution

Description

Density, distribution function, quantile function and random generation for the truncated binomial distribution.

Usage

```
dtbinom(x, size, prob, a = -Inf, b = Inf, log = FALSE)

ptbinom(q, size, prob, a = -Inf, b = Inf, lower.tail = TRUE, log.p = FALSE)

qtbinom(p, size, prob, a = -Inf, b = Inf, lower.tail = TRUE, log.p = FALSE)

rtbinom(n, size, prob, a = -Inf, b = Inf)
```

Arguments

x, q	vector of quantiles.
size	number of trials (zero or more).
prob	probability of success on each trial.
a, b	lower and upper truncation points ($a < x \leq b$).
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Examples

```

x <- rtbinom(1e5, 100, 0.83, 76, 86)
xx <- seq(0, 100)
plot(prop.table(table(x)))
lines(xx, dtbinom(xx, 100, 0.83, 76, 86), col = "red")
hist(ptbinom(x, 100, 0.83, 76, 86))

xx <- seq(0, 100, by = 0.01)
plot(ecdf(x))
lines(xx, ptbinom(xx, 100, 0.83, 76, 86), col = "red", lwd = 2)
uu <- seq(0, 1, by = 0.001)
lines(qtbinom(uu, 100, 0.83, 76, 86), uu, col = "blue", lty = 2)

```

TruncNormal

Truncated normal distribution

Description

Density, distribution function, quantile function and random generation for the truncated normal distribution.

Usage

```
dtnorm(x, mean = 0, sd = 1, a = -Inf, b = Inf, log = FALSE)
```

```

ptnorm(
  q,
  mean = 0,
  sd = 1,
  a = -Inf,
  b = Inf,
  lower.tail = TRUE,
  log.p = FALSE
)

```

```

qtnorm(
  p,
  mean = 0,
  sd = 1,
  a = -Inf,
  b = Inf,
  lower.tail = TRUE,
  log.p = FALSE
)

```

```
rtnorm(n, mean = 0, sd = 1, a = -Inf, b = Inf)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>mean, sd</code>	location and scale parameters. Scale must be positive.
<code>a, b</code>	lower and upper truncation points ($a < x \leq b$, with $a = -\text{Inf}$ and $b = \text{Inf}$ by default).
<code>log, log.p</code>	logical; if TRUE, probabilities p are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \frac{\phi\left(\frac{x-\mu}{\sigma}\right)}{\Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)}$$

Cumulative distribution function

$$F(x) = \frac{\Phi\left(\frac{x-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)}{\Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)}$$

Quantile function

$$F^{-1}(p) = \Phi^{-1}\left(\Phi\left(\frac{a-\mu}{\sigma}\right) + p \times \left[\Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)\right]\right)$$

For random generation algorithm described by Robert (1995) is used.

References

Robert, C.P. (1995). Simulation of truncated normal variables. *Statistics and Computing* 5(2): 121-125. <https://arxiv.org/abs/0907.4010>

Burkardt, J. (17 October 2014). The Truncated Normal Distribution. Florida State University. https://people.sc.fsu.edu/~jburkardt/presentations/truncated_normal.pdf

Examples

```
x <- rtnorm(1e5, 5, 3, b = 7)
hist(x, 100, freq = FALSE)
curve(dtnorm(x, 5, 3, b = 7), -8, 8, col = "red", add = TRUE)
hist(ptnorm(x, 5, 3, b = 7))
plot(ecdf(x))
curve(ptnorm(x, 5, 3, b = 7), -8, 8, col = "red", lwd = 2, add = TRUE)

R <- 1e5
partmp <- par(mfrow = c(2,4), mar = c(2,2,2,2))
```

```

hist(rtnorm(R), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x), -5, 5, col = "red", add = TRUE)

hist(rtnorm(R, a = 0), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, a = 0), -1, 5, col = "red", add = TRUE)

hist(rtnorm(R, b = 0), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, b = 0), -5, 5, col = "red", add = TRUE)

hist(rtnorm(R, a = 0, b = 1), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, a = 0, b = 1), -1, 2, col = "red", add = TRUE)

hist(rtnorm(R, a = -1, b = 0), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, a = -1, b = 0), -2, 2, col = "red", add = TRUE)

hist(rtnorm(R, mean = -6, a = 0), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, mean = -6, a = 0), -2, 1, col = "red", add = TRUE)

hist(rtnorm(R, mean = 8, b = 0), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, mean = 8, b = 0), -2, 1, col = "red", add = TRUE)

hist(rtnorm(R, a = 3, b = 5), freq= FALSE, main = "", xlab = "", ylab = "")
curve(dtnorm(x, a = 3, b = 5), 2, 5, col = "red", add = TRUE)

par(partmp)

```

TruncPoisson

Truncated Poisson distribution

Description

Density, distribution function, quantile function and random generation for the truncated Poisson distribution.

Usage

```

dtpois(x, lambda, a = -Inf, b = Inf, log = FALSE)

ptpois(q, lambda, a = -Inf, b = Inf, lower.tail = TRUE, log.p = FALSE)

qtpois(p, lambda, a = -Inf, b = Inf, lower.tail = TRUE, log.p = FALSE)

rtpois(n, lambda, a = -Inf, b = Inf)

```

Arguments

x, q vector of quantiles.
lambda vector of (non-negative) means.

<code>a, b</code>	lower and upper truncation points ($a < x \leq b$).
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as $\log(p)$.
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

References

Plackett, R.L. (1953). The truncated Poisson distribution. *Biometrics*, 9(4), 485-488.

Singh, J. (1978). A characterization of positive Poisson distribution and its statistical application. *SIAM Journal on Applied Mathematics*, 34(3), 545-548.

Dalgaard, P. (May 1, 2005). [R] simulate zero-truncated Poisson distribution. R-help mailing list. <https://stat.ethz.ch/pipermail/r-help/2005-May/070680.html>

Examples

```
x <- rtpois(1e5, 14, 16)
xx <- seq(-1, 50)
plot(prop.table(table(x)))
lines(xx, dtpois(xx, 14, 16), col = "red")
hist(ptpois(x, 14, 16))

xx <- seq(0, 50, by = 0.01)
plot(ecdf(x))
lines(xx, ptpois(xx, 14, 16), col = "red", lwd = 2)

uu <- seq(0, 1, by = 0.001)
lines(qtpois(uu, 14, 16), uu, col = "blue", lty = 2)

# Zero-truncated Poisson

x <- rtpois(1e5, 5, 0)
xx <- seq(-1, 50)
plot(prop.table(table(x)))
lines(xx, dtpois(xx, 5, 0), col = "red")
hist(ptpois(x, 5, 0))

xx <- seq(0, 50, by = 0.01)
plot(ecdf(x))
lines(xx, ptpois(xx, 5, 0), col = "red", lwd = 2)
lines(qtpois(uu, 5, 0), uu, col = "blue", lty = 2)
```

 TukeyLambda

Tukey lambda distribution

Description

Quantile function, and random generation for the Tukey lambda distribution.

Usage

```
qtlambda(p, lambda, lower.tail = TRUE, log.p = FALSE)
```

```
rtlambda(n, lambda)
```

Arguments

p	vector of probabilities.
lambda	shape parameter.
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
log.p	logical; if TRUE, probabilities p are given as $\log(p)$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Tukey lambda distribution is a continuous probability distribution defined in terms of its quantile function. It is typically used to identify other distributions.

Quantile function:

$$F^{-1}(p) = \begin{cases} \frac{1}{\lambda}[p^\lambda - (1-p)^\lambda] & \lambda \neq 0 \\ \log\left(\frac{p}{1-p}\right) & \lambda = 0 \end{cases}$$

References

Joiner, B.L., & Rosenblatt, J.R. (1971). Some properties of the range in samples from Tukey's symmetric lambda distributions. *Journal of the American Statistical Association*, 66(334), 394-399.

Hastings Jr, C., Mosteller, F., Tukey, J.W., & Winsor, C.P. (1947). Low moments for small samples: a comparative study of order statistics. *The Annals of Mathematical Statistics*, 413-426.

Examples

```
pp = seq(0, 1, by = 0.001)
partmp <- par(mfrow = c(2,3))
plot(qtlambda(pp, -1), pp, type = "l", main = "lambda = -1 (Cauchy)")
plot(qtlambda(pp, 0), pp, type = "l", main = "lambda = 0 (logistic)")
plot(qtlambda(pp, 0.14), pp, type = "l", main = "lambda = 0.14 (normal)")
plot(qtlambda(pp, 0.5), pp, type = "l", main = "lambda = 0.5 (concave)")
```

```

plot(qtlambda(pp, 1), pp, type = "l", main = "lambda = 1 (uniform)")
plot(qtlambda(pp, 2), pp, type = "l", main = "lambda = 2 (uniform)")

hist(rtlambda(1e5, -1), freq = FALSE, main = "lambda = -1 (Cauchy)")
hist(rtlambda(1e5, 0), freq = FALSE, main = "lambda = 0 (logistic)")
hist(rtlambda(1e5, 0.14), freq = FALSE, main = "lambda = 0.14 (normal)")
hist(rtlambda(1e5, 0.5), freq = FALSE, main = "lambda = 0.5 (concave)")
hist(rtlambda(1e5, 1), freq = FALSE, main = "lambda = 1 (uniform)")
hist(rtlambda(1e5, 2), freq = FALSE, main = "lambda = 2 (uniform)")
par(partmp)

```

Wald

Wald (inverse Gaussian) distribution

Description

Density, distribution function and random generation for the Wald distribution.

Usage

```

dwald(x, mu, lambda, log = FALSE)

pwald(q, mu, lambda, lower.tail = TRUE, log.p = FALSE)

rwald(n, mu, lambda)

```

Arguments

x, q	vector of quantiles.
mu, lambda	location and shape parameters. Scale must be positive.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.
p	vector of probabilities.

Details

Probability density function

$$f(x) = \sqrt{\frac{\lambda}{2\pi x^3}} \exp\left(\frac{-\lambda(x - \mu)^2}{2\mu^2 x}\right)$$

Cumulative distribution function

$$F(x) = \Phi\left(\sqrt{\frac{\lambda}{x}}\left(\frac{x}{\mu} - 1\right)\right) + \exp\left(\frac{2\lambda}{\mu}\right) \Phi\left(\sqrt{\frac{\lambda}{x}}\left(\frac{x}{\mu} + 1\right)\right)$$

Random generation is done using the algorithm described by Michael, Schucany and Haas (1976).

References

Michael, J.R., Schucany, W.R., and Haas, R.W. (1976). Generating Random Variates Using Transformations with Multiple Roots. *The American Statistician*, 30(2): 88-90.

Examples

```
x <- rwald(1e5, 5, 16)
hist(x, 100, freq = FALSE)
curve(dwald(x, 5, 16), 0, 50, col = "red", add = TRUE)
hist(pwald(x, 5, 16))
plot(ecdf(x))
curve(pwald(x, 5, 16), 0, 50, col = "red", lwd = 2, add = TRUE)
```

ZIB

Zero-inflated binomial distribution

Description

Probability mass function and random generation for the zero-inflated binomial distribution.

Usage

```
dzib(x, size, prob, pi, log = FALSE)

pzib(q, size, prob, pi, lower.tail = TRUE, log.p = FALSE)

qzib(p, size, prob, pi, lower.tail = TRUE, log.p = FALSE)

rzib(n, size, prob, pi)
```

Arguments

x, q	vector of quantiles.
size	number of trials (zero or more).
prob	probability of success in each trial. $0 < \text{prob} \leq 1$.
pi	probability of extra zeros.
log, log.p	logical; if TRUE, probabilities p are given as log(p).
lower.tail	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
p	vector of probabilities.
n	number of observations. If $\text{length}(n) > 1$, the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \pi + (1 - \pi)(1 - p)^n & x = 0 \\ (1 - \pi) \binom{n}{x} p^x (1 - p)^{n-x} & x > 0 \end{cases}$$

See Also

[Binomial](#)

Examples

```
x <- rzib(1e5, 10, 0.6, 0.33)
xx <- -2:20
plot(prop.table(table(x)), type = "h")
lines(xx, dzib(xx, 10, 0.6, 0.33), col = "red")

xx <- seq(0, 20, by = 0.01)
plot(ecdf(x))
lines(xx, pzib(xx, 10, 0.6, 0.33), col = "red")
```

ZINB

Zero-inflated negative binomial distribution

Description

Probability mass function and random generation for the zero-inflated negative binomial distribution.

Usage

```
dzinb(x, size, prob, pi, log = FALSE)

pzinb(q, size, prob, pi, lower.tail = TRUE, log.p = FALSE)

qzinb(p, size, prob, pi, lower.tail = TRUE, log.p = FALSE)

rzinb(n, size, prob, pi)
```

Arguments

x, q	vector of quantiles.
size	target for number of successful trials, or dispersion parameter (the shape parameter of the gamma mixing distribution). Must be strictly positive, need not be integer.
prob	probability of success in each trial. $0 < \text{prob} \leq 1$.

<code>pi</code>	probability of extra zeros.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as <code>log(p)</code> .
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \pi + (1 - \pi)p^r & x = 0 \\ (1 - \pi) \binom{x+r-1}{x} p^r (1-p)^x & x > 0 \end{cases}$$

See Also

[NegBinomial](#)

Examples

```
x <- rzinb(1e5, 100, 0.6, 0.33)
xx <- -2:200
plot(prop.table(table(x)), type = "h")
lines(xx, dzinb(xx, 100, 0.6, 0.33), col = "red")

xx <- seq(0, 200, by = 0.01)
plot(ecdf(x))
lines(xx, pzinb(xx, 100, 0.6, 0.33), col = "red")
```

ZIP

Zero-inflated Poisson distribution

Description

Probability mass function and random generation for the zero-inflated Poisson distribution.

Usage

```
dzip(x, lambda, pi, log = FALSE)

pzip(q, lambda, pi, lower.tail = TRUE, log.p = FALSE)

qzip(p, lambda, pi, lower.tail = TRUE, log.p = FALSE)

rzip(n, lambda, pi)
```

Arguments

<code>x, q</code>	vector of quantiles.
<code>lambda</code>	vector of (non-negative) means.
<code>pi</code>	probability of extra zeros.
<code>log, log.p</code>	logical; if TRUE, probabilities <code>p</code> are given as <code>log(p)</code> .
<code>lower.tail</code>	logical; if TRUE (default), probabilities are $P[X \leq x]$ otherwise, $P[X > x]$.
<code>p</code>	vector of probabilities.
<code>n</code>	number of observations. If <code>length(n) > 1</code> , the length is taken to be the number required.

Details

Probability density function

$$f(x) = \begin{cases} \pi + (1 - \pi)e^{-\lambda} & x = 0 \\ (1 - \pi)\frac{\lambda^x e^{-\lambda}}{x!} & x > 0 \end{cases}$$

See Also

[Poisson](#)

Examples

```
x <- rzip(1e5, 6, 0.33)
xx <- -2:20
plot(prop.table(table(x)), type = "h")
lines(xx, dzip(xx, 6, 0.33), col = "red")

xx <- seq(0, 20, by = 0.01)
plot(ecdf(x))
lines(xx, pzip(xx, 6, 0.33), col = "red")
```

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